## **Biomedical Applications and the Probabilistic Framework**

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## **Talk Overview**

- Motivation of Probabilistic Concepts
- BCI, current practice & shortcomings
- Probabilistic Kalman Filter
- Adaptive BCI
- Gene Discovery
- DAG for Bayesian Marker Identification
- Gene Selection
- Discussion of Model Selection

## **Probabilistic Motivations**



Thomas Bayes (1701 - 1763) Learning from data using a decision theoretic framework

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 $p(x|\mathcal{D}) = \frac{p(\mathcal{D}|x)p(x)}{p(\mathcal{D})}$ First consequence: we must revise beliefs according to Bayes theorem

## **Probabilistic Motivations**



Thomas Bayes (1701 - 1763) Learning from data using a decision theoretic framework

 $p(x|\mathcal{T})$ First cording to Bayes theorem pected utilities

 $\overline{\alpha_{opt}} = \operatorname{argmax}_{\alpha} < u(\alpha) >$ , where  $< u(\alpha) > = \int_x u(\alpha, x) p(x|\mathcal{D}) dx.$ consequence: we Second consequence: Demust revise beliefs ac- cisions by maximising ex-

## **Brain Computer Interface**



### Computer is controlled *directly* by *cortical activity*.

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## **Classification of BCIs**



intracranial EEG - > high spatial and temporal resolution; highly invasive!; allows 2-d control of artificial limb.

surface EEG - > low spatial and temporal resolution; no permanent interference with patient; slow! at most 20 bit per minute and task.

- > focus on BCI's based on scalp recordings.

– > low bit rates; last resort if no other communica tion possible

## **BCI with almost no adaptation**

- P300 based: L. A. Farwell and E. Donchin, -> User intention is embedded within a sequence of symbols. The correct symbol leads to "surprise" and triggers a P300.
- Filter & threshold: N. Birbaumer etal., -> threshold slow cortical potentials; J.R. Wolpaw etal., -> threshold moving average in an appropriate pass band e.g. μ-rhythm.

These principles rely mostly on user training.

## **BCI & static pattern recognition**

- Extract representation of EEG "waveforms" (e.g. low pass filtered time series; spectral representation)
- Parameterize supervised classification implicitly assuming stationarity.

### What if

Technical setup changes during operation? (e.g. electrolyte changes impedance) User learns from feedback? User shows fatigue? Assuming stationarity must be wrong !

– > Probabilistic method for "adaptive" BCI.





Key: get  $\lambda$  right (may regard  $1/\lambda$  as learning rate)



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Illustration of  $\langle \lambda \rangle$  and "instantaneous" generalization error for B. D. Ripley's synthetic data with artificial non-stationarity (swap labels after sample 500).

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## **Adaptive BCI**

### by variational Kalman filtering.

BCI: data driven prediction of cognitive state from EEG measurements. Working hypothesis: EEG dynamics during a cognitive task are subject to temporal variation (learning effects, fatigue ...) Represent EEG segments by z-transformed reflection coefficients.

Mutual information, adaptive method and identical 'stationary' model.



## **Communication Bandwidth**

	bit rates $r_{P(y)}$ [bit/s]		
task	vkf	vsi	$P_{null}$
rest/move no fb.	0.18	0.10	$\ll 0.01$
rest/move fb.	0.18	0.13	$\ll 0.01$
move/math no fb.	0.18	0.11	$\ll 0.01$
move/math fb.	0.15	0.10	$\ll 0.01$
nav./aud./move	0.55	0.49	$\ll 0.01$
audit./move	0.38	0.35	$\ll 0.01$
navig./move	0.32	0.28	$\ll 0.01$
navig./audit.	0.37	0.34	$\ll 0.01$

**Conclusion:** adaptive methods increase BCI band-widths even on short time scales.

## **Gene discovery**

Discovering "important" genes (or proteins) from microarray datasets can be classified as

- Identification of all differentially expressed genes.
- Identification of reliable (sets) of marker genes.

Current practise for the first: classical methods (e.g. t-test on differences of means) or probabilistic approaches with one indicator variable for each gene.

The second is typically done by conventional feature subset selection. As a result we obtain a set of genes that was found by heuristic search.

## **Bayesian Marker Identification**

Missing in FSS: How good are other explanations? Interpret microarray data as classification problem of "genetic" regressors w.r.t. a discrete response.

-> Bayesian variable selection provides this information. However: hopeless, unless we constrain the dimensionality.

Simplified attempt: - > Find distribution over individual genes.

Probabilities result from the marginal likelihood of each model.

$$P(I|\mathcal{D}) = \frac{\int_{\boldsymbol{w}} p(\mathcal{D}|\boldsymbol{w}) p(\boldsymbol{w}|I) P(I)}{\sum_{I} \int_{\boldsymbol{w}} p(\mathcal{D}|\boldsymbol{w}) p(\boldsymbol{w}|I) P(I) d\boldsymbol{w}},$$

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## **DAG for Marker Identification**



Latent variable probit GLM.

 $z_n$  is a one dimensional Gaussian random variable with mean  $w^T x_n$  and precision 1.

 $P(y_n \equiv 1 | z_n) = \begin{cases} 1, \text{ if } z_n > 0\\ 0, \text{ if } z_n \le 0 \end{cases}$ 

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Inference can be done by a variational method (systematic error) or by sampling (random error). The latter allows to integrate over w analytically and we draw from  $z_n$  and I only.

## **Asymptotic Behaviour**



## **Comparison with ML**



Results differ since Bayesian model posteriors take "complexity" (ref. Hochreiter's "flat minima") into account.

## Selection and Gen. Accuracy

# Most probable regressors selected at a 0.99 threshold Generalization accuracy

Acc. no.	description	$P(I \mathcal{D})$	
Colon Cancer (Alon et. al.)			
Z50753	Uroguanylin	0.76	
R87126	Myosin	0.21	
M63391	desmin gene	0.01	
M36634	vasoact. pept.	0.01	
Leukaemia (Golub et. al.)			
X95735	Zyxin	0.93	
M55150	FAH Fumarylac.	0.05	
M27891	CST3 Cystatin C	0.01	

Dataset	B. probit	"indifference"		
Colon	84%	74% to 94%		
Leukaemia	88%	91% to 96%		
No 'better' results in literature $- >$				
confirms model.				
Biology confirms Uroguanylin (cell				
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Biology confirms Uroguanylin (cell				
apoptosis) as important in colon can-				
cer development.				
But: Meaning of the prob-				
abilities?				

## Discussion

Quoting  $P(I|\mathcal{D}) \rightarrow \mathcal{M}$ -closed model selection with zero-one utility.

Our approach should assume an  $\mathcal{M}$ -open scenario. Under asymptotic normality,  $P(I|\mathcal{D})$  degenerates on  $I_i \in \mathcal{M}$  that minimizes  $\int p(y|\boldsymbol{w}_t) \log(p(y|\hat{\boldsymbol{w}}_i)/p(y|\boldsymbol{w}_t)) dy).$ 

If the predictive distribution of a new observation is of interest, B&S's suggest to use a logarithmic score function for  $\mathcal{M}$ -open model comparison.

 $\int \log(p(y|I_i, \mathcal{D}))p(y|\mathcal{D})dy$ 

(e.g. cross validation estimate, still to be done)

## A simple idea:

the world is one probabilistic model.

- Applications often require hierarchical structure: a feature extraction part and a probabilistic model.
- Classical approach: treat both parts separately and thus regard features as sufficient statistic of the data. — > Features are deterministic variables.
- Our suggestion: treat such hierarchical settings as one probabilistic model. — > Feature extraction is a representation in a latent space.

## **Bayes' Consistent Models**



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Expected utility requires to integrate over all unknown variables, including  $\varphi_a$ ,  $\varphi_b$ ,  $I_a$  and  $I_b$  that represent a feature space.

## **Bayes' Consistent Models**





**Expected utility** requires to integrate over all unknown variables, includ- Decisions depend on ing  $\varphi_a$ ,  $\varphi_b$ ,  $I_a$  and  $I_b$  that (un)certainty and may represent a feature space. thus change.

## **Time Series Classification**

### **ROC Curves**



### Kullback Leibler Divergence



## **More Results**

### Expected feature values



### Kullback Leibler Divergence for "Artefacts"



## **Variational Kalman Filter**

The logarithmic model evidence for a window of size N is

$$\log(p(\mathcal{D}_N)) = \log\left(\int_{\lambda} \prod_{n=1}^{N} \left[\int_{\boldsymbol{w}_{n-1}} \int_{\boldsymbol{w}_n} p(\boldsymbol{w}_{n-1} | \mathcal{D}_{n-1}) \right. \\ \left. p(\boldsymbol{w}_n | \boldsymbol{w}_{n-1}, \lambda \boldsymbol{I}) P(y_n | \boldsymbol{w}_n, \boldsymbol{\phi}_n) d\boldsymbol{w}_n d\boldsymbol{w}_{n-1} \right] p(\lambda | \alpha, \beta) d\lambda \right).$$

This is not a probabilistic structure! (need Rauch Tung Striebel smoother)

## **Variational Kalman Filter**

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$$p(\boldsymbol{w}_n | \boldsymbol{w}_{n-1}, \lambda \boldsymbol{I}) P(y_n | \boldsymbol{w}_n, \boldsymbol{\phi}_n) d\boldsymbol{w}_n d\boldsymbol{w}_{n-1} \right] p(\lambda | \alpha, \beta) d\lambda.$$

This is not a probabilistic structure! (need Rauch Tung Striebel smoother) Plug in distributions and integrate over  $w_{n-1}$ :

$$\log(p(\mathcal{D}_N)) = \log\left(\int_{\lambda} \prod_{n=1}^{N} \left[\int_{\boldsymbol{w}_n} (2\pi)^{-\frac{d}{2}} |\boldsymbol{\Lambda}_{n-1}^{-1} + \lambda^{-1}\boldsymbol{I}|^{-\frac{1}{2}} \right] \times \exp(-0.5(\boldsymbol{w}_n - \boldsymbol{\hat{w}}_{n-1})^T (\boldsymbol{\Lambda}_{n-1}^{-1} + \lambda^{-1}\boldsymbol{I})^{-1}(\boldsymbol{w}_n - \boldsymbol{\hat{w}}_{n-1})) \times (1 + \exp((2\boldsymbol{y}_n - 1)\boldsymbol{\phi}_n^T \boldsymbol{w}_n))^{-1} d\boldsymbol{w}_n \right] \times \frac{\beta^{\alpha}}{\Gamma(\alpha)} \lambda^{(\alpha-1)} \exp(-\beta\lambda) d\lambda$$

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## **Lower Bounds**

$$\log(P(y_n|\boldsymbol{\phi}_n, \boldsymbol{w}_n)) \geq -\frac{(2y_n - 1)\boldsymbol{\phi}_n^T \boldsymbol{w}_n}{2} - \log(2) - \log(\cosh(\frac{\xi_n}{2})) \\ - \frac{\tanh(\frac{\xi_n}{2})}{4\xi_n} \left( \left(\frac{\boldsymbol{\phi}_n^T \boldsymbol{w}_n}{2}\right)^2 - \xi_n^2 \right)$$

#### back to vkf

## **Lower Bounds**

$$\begin{split} \log(P(y_n|\boldsymbol{\phi}_n, \boldsymbol{w}_n)) &\geq -\frac{(2y_n - 1)\boldsymbol{\phi}_n^T \boldsymbol{w}_n}{2} - \log(2) - \log(\cosh(\frac{\xi_n}{2})) \\ &- \frac{\tanh(\frac{\xi_n}{2})}{4\xi_n} \left( \left(\frac{\boldsymbol{\phi}_n^T \boldsymbol{w}_n}{2}\right)^2 - \xi_n^2 \right) \\ &- 0.5 \log|\boldsymbol{\Lambda}_{n-1}^{-1} + \lambda^{-1}\boldsymbol{I}| \geq \frac{d}{2} \log \lambda - \frac{1}{2} \log|\boldsymbol{\nu}\boldsymbol{\Lambda}_n^{-1} + \boldsymbol{I}| \\ &- \frac{1}{2} (\lambda - \boldsymbol{\nu}) \operatorname{tr}(\boldsymbol{\nu}\boldsymbol{I} + \boldsymbol{\Lambda}_n)^{-1}, \end{split}$$

back to vkf

## **Lower Bounds**

$$\log(P(y_{n}|\phi_{n}, w_{n})) \geq -\frac{(2y_{n}-1)\phi_{n}^{T}w_{n}}{2} - \log(2) - \log(\cosh(\frac{\xi_{n}}{2}))$$

$$- \frac{\tanh(\frac{\xi_{n}}{2})}{4\xi_{n}} \left( \left(\frac{\phi_{n}^{T}w_{n}}{2}\right)^{2} - \xi_{n}^{2} \right)$$

$$- 0.5 \log|\mathbf{\Lambda}_{n-1}^{-1} + \lambda^{-1}\mathbf{I}| \geq \frac{d}{2}\log\lambda - \frac{1}{2}\log|\nu\mathbf{\Lambda}_{n}^{-1} + \mathbf{I}|$$

$$- \frac{1}{2}(\lambda - \nu)\operatorname{tr}(\nu\mathbf{I} + \mathbf{\Lambda}_{n})^{-1},$$

$$- 0.5(w_{n} - \hat{w}_{n-1})^{T}(\mathbf{\Lambda}_{n-1}^{-1} + \lambda^{-1}\mathbf{I})^{-1}(w_{n} - \hat{w}_{n-1}) \geq$$

$$- 0.5(w_{n} - \hat{w}_{n-1})^{T}(\mathbf{\Lambda}_{n-1}^{-1} + \nu^{-1}\mathbf{I})^{-1}(w_{n} - \hat{w}_{n-1})$$

$$- 0.5(\lambda - \nu)(w_{n} - \hat{w}_{n-1})^{T}(\nu\mathbf{\Lambda}_{n-1}^{-1} + \mathbf{I})^{-2}(w_{n} - \hat{w}_{n-1})$$

back to vkf